MULTILOG Example #5

SUDAAN Statements and Results Illustrated

- PREDMARG
- ADJRR option
- GENLOGIT option
- PRED_EFF
- SUBPOPX

Input Data Set(s): NH3MI1.SAS7bdat - NH3MI5.SAS7bdat

Example

Using data from the NHANES III Multiply Imputed (MI) dataset, model the effects of age, weight, gender, and current smoking status among adults age 20 yrs and older on the self-rating of health status.

In addition to featuring analysis of multiply imputed data, this example also highlights the estimation of model-adjusted risks, risk ratios, and risk differences for a multinomial outcome via predicted marginal proportions (ADJRR option on PREDMARG statement). Confidence intervals for model-adjusted risks (marginals) are new in Release 11.0.

Solution

This example illustrates the MULTILOG SUDAAN procedure with the following three multiply imputed variables.

- HAB1MI = Self-rating of health status. This is a five-level categorical variable coded 1=Excellent, 2=Very good, 3=Good, 4=Fair, and 5=Poor.
- HAM6MI = Weight in pounds (continuous).
- HAR3RMI = Current smoking status (1=yes, 2=no).

The model also uses AGE (continuous, in years) and HSSEX (1=Male, 2=Female) as independent variables. Only one version of these variables is available. These two variables, along with all design statement variables, are repeated in *each* of the five MI datasets.

Five different versions of these imputation-revised variables were constructed and saved on the *Nh3mi1,...,Nh3mi5* SAS datasets, respectively. In the DATA option on the PROC MULTILOG statement, the user only needs to specify the *first* dataset in the series of five. The MI_COUNT=5 option tells SUDAAN that there are five MI datasets available for analysis, numbered sequentially.

The NEST statement indicates that SDPSTRA6 is the stratification variable and SDPPSU6 is the primary sampling unit. The WEIGHT statement indicates that WTPFQX6 contains the analysis weight for each record on the file.

The SUBPOPX statement restricts the analysis to adults aged 20 years or older.

In MULTILOG, the CLASS statement contains the dependent variable *and* all covariates that are to be modeled as categorical (HAB1MI is the dependent; HSSEX and HAR3RMI are the categorical independent).

The GENLOGIT (generalized logit) link specifies the generalized logit model. The default Wald-*F* test is used for all tests of hypotheses.

The PREDMARG statement requests the predicted marginal proportion (*model-adjusted risk*) for each level of HSSEX. The ADJRR option on the PREDMARG statement computes the *model-adjusted risk ratio* for males vs. females. Finally, the PRED_EFF statement requests the *difference* in predicted marginal proportions (*risk differences*) for Males vs. Females.

We include multiple PRINT statements, all of which are optional. Multiple PRINT statements allow the user to set up different default print environments (SETENV statements) for different PRINT groups. The PRINT statements are used in this example to request the PRINT groups of interest; to calculate individual statistics of interest, and in some cases, change default labels for those statistics; and to specify a variety of formats for those printed statistics. Without the PRINT statement, default statistics are produced from each PRINT group, with default formats.

The SETENV statements are optional. They set up default formats for printed statistics and further manipulate the printout to the needs of the user.

The RFORMAT and RLABEL statements associate SAS formats and variable labels, respectively, with the variables used in the procedure. Without the RLABEL statement, SAS variable labels would be produced if already defined.

This example was run in SAS-Callable SUDAAN, and the SAS program and *.LST files are provided.

Exhibit 1. SAS-Callable SUDAAN Code

```
libname in "c:\903winbetatest\nhanes3";
options nocenter linesize=85 pagesize=68;
proc format;
 value hab 1="1=Excellent"
            2="2=Very Good"
            3="3=Good"
            4="4=Fair"
            5="5=Poor";
 value sex 1="1=Male"
            2="2=Female";
 value smoke 1="1=Yes"
             2="2=No";
data mi1; set in.nh3mi1;
  AGE = HSAGEIR;
   IF HSAGEU = 1 THEN AGE = AGE / 12;
   keep sdpstra6 sdppsu6 wtpfqx6 age hssex hab1mi ham6mi har3rmi;
proc sort data=mi1; by SDPSTRA6 SDPPSU6;
data mi2; set in.nh3mi2;
  AGE = HSAGEIR;
   IF HSAGEU = 1 THEN AGE = AGE / 12;
   keep sdpstra6 sdppsu6 wtpfqx6 age hssex hab1mi ham6mi har3rmi;
proc sort data=mi2; by SDPSTRA6 SDPPSU6;
data mi3; set in.nh3mi3;
   AGE = HSAGEIR;
    IF HSAGEU = 1 THEN AGE = AGE / 12;
   keep sdpstra6 sdppsu6 wtpfqx6 age hssex hab1mi ham6mi har3rmi;
proc sort data=mi3; by SDPSTRA6 SDPPSU6;
data mi4; set in.nh3mi4;
     AGE = HSAGEIR;
     IF HSAGEU = 1 THEN AGE = AGE / 12;
     keep sdpstra6 sdppsu6 wtpfqx6 age hssex hab1mi ham6mi har3rmi;
proc sort data=mi4; by SDPSTRA6 SDPPSU6;
data mi5; set in.nh3mi5;
      AGE = HSAGEIR;
      IF HSAGEU = 1 THEN AGE = AGE / 12;
      keep sdpstra6 sdppsu6 wtpfqx6 age hssex hab1mi ham6mi har3rmi;
proc sort data=mi5; by SDPSTRA6 SDPPSU6;
```

Exhibit 1. SAS-Callable SUDAAN Code-cont.

```
PROC MULTILOG DATA=mi1 FILETYPE=SAS mi count=5 DESIGN=WR;
 NEST SDPSTRA6 SDPPSU6 / MISSUNIT;
 WEIGHT WTPFQX6;
 SUBPOPX age>=20 / name="Age 20+";
 CLASS HAB1MI HSSEX HAR3RMI;
 MODEL HAB1MI = AGE HAM6MI HSSEX HAR3RMI / genlogit;
 PREDMARG HSSEX / adjrr;
 PRED EFF HSSEX=(1 -1) / name="Males-Females";
 setenv labwidth=25 colspce=1 topmgn=0 colwidth=7 decwidth=4;
 print / betas=default risk=default tests=default t_betafmt=f6.2 waldffmt=f6.2
         waldpfmt=f7.4 dffmt=f7.0 orfmt=f10.3 loworfmt=f9.3 uporfmt=f9.3
         style=nchs;
 SETENV LABWIDTH=25 COLWIDTH=7 DECWIDTH=4 COLSPCE=3 TOPMGN=0;
 PRINT / PRED MRG=default PRMGCONS=default lowpmfmt=f6.4 uppmfmt=f6.4
         predmrgfmt=f9.4 t prdmrgfmt=f8.2 prmgconfmt=f8.4 t pmconfmt=f7.2;
 SETENV LABWIDTH=30 COLWIDTH=5 DECWIDTH=3 COLSPCE=2 TOPMGN=0;
 PRINT PRED RR="Risk Ratio" / PREDRISK=default;
 RLABEL HAB1MI="Self-Rating Health Status";
 RLABEL HAM6MI="Weight (lbs)";
 RLABEL HAR3RMI="Current Smoker?";
 RFORMAT hab1mi hab.;
 RFORMAT hssex sex.;
 RFORMAT har3rmi smoke.;
 RTITLE "Modelling Self-Rating of Health Status (MI Data)";
 RFOOTNOTE "NHANES III, Multiply Imputed Dataset: Respondents Age 20+";
```

Exhibit 2. First Page of SUDAAN Output

SUDAAN Software for the Statistical Analysis of Correlated Data Copyright Research Triangle Institute November 2011 Release 11.0.0 DESIGN SUMMARY: Variances will be computed using the Taylor Linearization Method, Assuming a With Replacement (WR) Design Sample Weight: WTPFQX6 Stratification Variables(s): SDPSTRA6 Primary Sampling Unit: SDPPSU6 Processing data for set 1 of imputed variables: Processing data for set 2 of imputed variables: Processing data for set 3 of imputed variables: Processing data for set 4 of imputed variables: Processing data for set 5 of imputed variables: Processing data for set 1 of imputed variables: Independence parameters have converged in 6 iterations : 33994 Number of observations read Weighted count:251097002 Observations in subpopulation : 18825 Weighted count:177180670 Weighted count:177180670 Observations used in the analysis : 18825 Denominator degrees of freedom 49 : Maximum number of estimable parameters for the model is 20 File MI1 contains 98 Clusters 98 clusters were used to fit the model Maximum cluster size is 297 records Minimum cluster size is 75 records Sample and Population Counts for Response Variable HAB1MI Based on observations used in the analysis 1=Excellent:Sample Count2824Population Count367947942=Very Good:Sample Count4387Population Count540820743=Good:Sample Count6743Population Count575988504=Fair:Sample Count3833Population Count22784573 5=Poor : Sample Count 1038 Population Count 5920379

Exhibit 2. First Page of SUDAAN Output-cont.

Processing data for set 2 of imputed variables: Independence parameters have converged in 6 iterations Number of observations read : 33994 Weighted count:251097002 Observations in subpopulation : 18825 Weighted count:177180670 Observations used in the analysis : 18825 Weighted count:177180670 Denominator degrees of freedom : 49 Maximum number of estimable parameters for the model is 20 File MI2 contains 98 Clusters 98 clusters were used to fit the model Maximum cluster size is 297 records Minimum cluster size is 75 records Sample and Population Counts for Response Variable HAB1MI Based on observations used in the analysis 1=Excellent:Sample Count2823Population Count367897932=Very Good:Sample Count4389Population Count541140953=Good:Sample Count6743Population Count576097984=Fair:Sample Count3833Population Count227579285=Poor:Sample Count1037Population Count5909056 Processing data for set 3 of imputed variables: Independence parameters have converged in 6 iterations Number of observations read : 33994 Observations in subpopulation : 18825 Weighted count:251097002 Weighted count:177180670 Observations used in the analysis : 18825 Weighted count:177180670 Denominator degrees of freedom : 49 Maximum number of estimable parameters for the model is 20 File MI3 contains 98 Clusters 98 clusters were used to fit the model Maximum cluster size is 297 records Minimum cluster size is 75 records Sample and Population Counts for Response Variable HAB1MI Based on observations used in the analysis 1=Excellent: Sample Count 2823 Population Count 36789793 2=Very Good: Sample Count 4390 Population Count 54095547 3=Good : Sample Count 6740 Population Count 57614324 4=Fair : Sample Count 3835 Population Count 22771679 : Sample Count 1037 5=Poor Population Count 5909327

Exhibit 2. First Page of SUDAAN Output-cont.

```
Processing data for set 4 of imputed variables:
Independence parameters have converged in 6 iterations
Number of observations read : 33994
Observations in subpopulation : 18825
Number of observations read
                                        : 33994
                                                       Weighted count:251097002
                                                       Weighted count:177180670
Observations used in the analysis : 18825
                                                       Weighted count:177180670
Denominator degrees of freedom :
                                              49
Maximum number of estimable parameters for the model is 20
File MI4 contains 98 Clusters
 98 clusters were used to fit the model
Maximum cluster size is 297 records
Minimum cluster size is 75 records
Sample and Population Counts for Response Variable HAB1MI
Based on observations used in the analysis
 1=Excellent:Sample Count2824Population Count367910512=Very Good:Sample Count4388Population Count540956073=Good:Sample Count6738Population Count57585236
  4=Fair
           : Sample Count 3837 Population Count 22788126
: Sample Count 1038 Population Count 5920649
  5=Poor
Processing data for set 5 of imputed variables:
Independence parameters have converged in 6 iterations
Number of observations read : 33994
Observations in subpopulation : 18825
                                                       Weighted count:251097002
                                                      Weighted count:177180670
Observations used in the analysis : 18825
                                                       Weighted count:177180670
Denominator degrees of freedom :
                                               49
Maximum number of estimable parameters for the model is 20
File MI5 contains 98 Clusters
  98 clusters were used to fit the model
Maximum cluster size is 297 records
Minimum cluster size is 75 records
Sample and Population Counts for Response Variable HAB1MI
Based on observations used in the analysis
  1=Excellent: Sample Count 2824 Population Count 36808280
 2=Very Good:Sample Count4387Population Count540826973=Good:Sample Count6741Population Count540826974=Fair:Sample Count6741Population Count576022955=Poor:Sample Count3836Population Count227783415=Poor:Sample Count1037Population Count5909056
Overall degrees of freedom (Rubin): 46.36
-2 * Normalized Log-Likelihood with Intercepts Only : 53880.98
-2 * Normalized Log-Likelihood Full Model : 51917.42
Approximate Chi-Square (-2 * Log-L Ratio) : 1963.56
                                                              : 1963.56
Degrees of Freedom
                                                              •
                                                                        16
Note: The approximate Chi-Square is not adjusted for clustering.
       Refer to hypothesis test table for adjusted test.
```

The sample design information is followed by summary information on each of the imputed datasets: convergence of parameters; observations read in subpopulation and used in analysis; number of clusters; minimum and maximum cluster size; and a distribution of the response variable based on observations used in the analysis (*Exhibit 2*). In each dataset, there were 33,994 observations read, with 18,825 used

in the analysis. The distribution of the response variable changed slightly from dataset to dataset, since HAB1MI is one of the multiply imputed variables in the analysis.

The following results represent the summary over all imputations (the option BY_MI on the PRINT statement produces results separately for each imputation). The frequency distributions for variables on the CLASS statement are presented first (*Exhibit 3* to *Exhibit 5*).

Exhibit 3. CLASS Variable Frequencies (HAB1MI)

```
Frequencies and Values for CLASS Variables
Results for Summary Over All Imputations
by: Self-Rating Health Status.
  _____
Self-Rating
Health
                         Value
Status
            Frequency
_____
Ordered
Position:
 1
             2823.600 1=Excellent
Ordered
 Position:
            4388.200 2=Very Good
 2
Ordered
 Position:
             6741.000 3=Good
 3
Ordered
Position:
             3834.800
                        4=Fair
 Δ
Ordered
 Position:
 5
             1037.400
                     5=Poor
_____
```

Exhibit 4. CLASS Variable Frequencies (Sex)

```
Frequencies and Values for CLASS Variables
Results for Summary Over All Imputations
by: Sex.
------
Sex
       Frequency Value
-----
Ordered
 Position:
            8816 1=Male
 1
Ordered
 Position:
            10009 2=Female
 2
------
```

Exhibit 5. CLASS Variable Frequencies (Current Smoker?)

Exhibit 6. Estimated Regression Coefficients

Variance Estimation Method: Taylor Series (WR) Using Multiply Imputed Data SE Method: Robust (Binder, 1983) Working Correlations: Independent Link Function: Generalized Logit Response variable HAB1MI: Self-Rating Health Status For Subpopulation: Age 20+

Modelling Self-Rating of Health Status (MI Data)

Results for Summary Over All Imputations

HAB1MI (log-odds) Independent Variables and Effects	Beta		Lower 95% Limit	Upper 95% Limit	T-Test	P-value T-Test
		SE Bela	веца	веса	B=0	в=0
l=Excellent vs 5=Poor						
Intercept	7.1332	0.4295	6.2691	7.9974	16.61	0.0000
Sex						
1=Male	0.6362	0.1125	0.4099	0.8625	5.65	0.0000
2=Female	0.0000	0.0000				
Current Smoker?						
1=Yes	-1.2645	0.1637	-1.5938	-0.9351	-7.72	0.0000
2=No	0.0000	0.0000				
AGE	-0.0701	0.0036	-0.0773	-0.0629	-19.60	0.0000
Weight (lbs)	-0.0104	0.0018	-0.0140	-0.0069	-5.90	0.0000
2=Very Good vs 5=Poor						
Intercept	6.4642	0.4335	5.5919	7.3365	14.91	0.0000
Sex	0.1012	0.1000	0.0010		± · • > ±	0.0000
1=Male	0.3354	0.1046	0.1251	0.5458	3.21	0.0024
2=Female	0.0000	0.0000	0.1201	0.0100		0.0021
Current Smoker?	0.0000	0.0000	•	•	•	•
1=Yes	-0 8281	0 1563	-1 1424	-0 5137	-5 30	0 0000
2=No	0.0201	0.1000	1.1121	0.0107	5.50	0.0000
ACE	-0.0658	0.0000	-0.0731	-0 0585	-18 05	
Meight (lbg)	-0.0047	0.0030	-0.0084	-0.0009	-2 52	0.0000
Record we 5-Roor	0.0047	0.0010	0.0004	0.0005	2.52	0.0152
Intercept	5 3203	0 3660	1 5030	6 0566	11 51	0 0000
Incercept	5.5205	0.3000	4.3039	0.0500	14.04	0.0000
Jewala	0 15/1	0 1120	0 0749	0 2020	1 25	0 1000
	0.1341	0.1130	-0.0748	0.3029	1.30	0.1022
2-remare	0.0000	0.0000	•	•	•	•
Laves	0 4204	0 1 5 7 4	0 7501	0 1007	0 70	0 0070
1=ies	-0.4394	0.1574	-0./561	-0.1227	-2.19	0.0076
Z=NO	0.0000	0.0000			17 07	•
AGE	-0.0536	0.0031	-0.0599	-0.04/4	-1/.2/	0.0000
Weight (LDS)	-0.0008	0.001/	-0.0042	0.0027	-0.45	0.65//
H=Fair VS S=Poor	0.0000	0 4000	0 0050	2 7 2 0 0	c 0.0	0 0000
Intercept	2.8866	0.4232	2.0352	3./380	6.82	0.0000
Sex						
l=Male	-0.1220	0.1292	-0.3819	0.1379	-0.94	0.3497
2=Female	0.0000	0.0000	•	•	•	•
Current Smoker?						
l=Yes	-0.1732	0.1515	-0.4780	0.1316	-1.14	0.2588
2=No	0.0000	0.0000	•	•	•	•
AGE	-0.0312	0.0034	-0.0380	-0.0244	-9.25	0.0000
Weight (lbs)	0.0020	0.0017	-0.0015	0.0054	1.15	0.2563

The generalized logit model fits a separate logit equation for *each* of the response categories vs. the last (*Exhibit 6*). Males have a significantly increased chance of being in the Excellent self-reported health status category compared to women. Being a non-smoker, as well as having lowered weight and age, are also significantly associated with self-reporting Excellent or Very Good health status.

Exhibit 7. ANOVA Table

Variance Estimation Method: Tay SE Method: Robust (Binder, 1983 Working Correlations: Independe Link Function: Generalized Logi Response variable HAB1MI: Self- For Subpopulation: Age 20+	lor Series (WF) nt t Rating Health	8) Using M Status	ultiply Imp	uted Da	ata	
Modelling Self-Rating of Health	Status (MI Da	ita)				
Results for Summary Over All Im	putations					
Contrast	Degrees of Freedom	Wald F	P-value Wald F			
OVERALL MODEL MODEL MINUS INTERCEPT INTERCEPT HSSEX HAR3RMI AGE HAM6MI	20 16 4 4 4 4 4	146.23 53.21	0.0000 0.0000 0.0000 0.0000 0.0000 0.0000			
NHANES III, Multiply Imputed Da	taset: Respon	idents Age	20+			

Exhibit 7 indicates that all four independent variables (sex, current smoker, age, and weight) are significantly associated with HAB1MI, self-rating of general health status among people 20 years of age and older. Note that each variable effect has four df, since there are four logit equations.

Exhibit 8. Default Odds Ratios

Variance Estimation Method SE Method: Robust (Binder, Working Correlations: Inde Link Function: Generalized Response variable HAB1MI:	d: Taylor Seri 1983) ependent d Logit Self-Rating H	es (WR) Usino ealth Status	g Multiply	Imputed Data
For Subpopulation: Age 20-	÷			
Modelling Self-Rating of H	Health Status	(MI Data)		
Results for Summary Over A	All Imputation	S		
HAB1MT (log-odds)				
Independent Variables and Effects	Odds Ratio	Lower 95% Limit OR	Upper 95% Limit OR	
1=Evcellent vs 5=Poor				
Intercept Sex	1252.933	527.997	2973.201	
1=Male 2=Female	1.889 1.000	1.507	2.369	
Current Smoker?				
1=Yes	0.282	0.203	0.393	
2=No	1.000	• • • • • •	• • • • •	
AGE Wojcht (lbs)	0.932	0.926	0.939	
Wery Good we 5-Poor	0.990	0.900	0.995	
Intercept	641.755	268.241	1535.372	
Sex	1 200	1 1 2 2	1 700	
1=Male	1.399	1.133	1.726	
Z=Female	1.000	•	•	
1-Vee	0 /37	0 319	0 598	
2=No	1 000	0.319	0.590	
AGE	0.936	0.929	0.943	
Weight (lbs)	0.995	0.992	0.999	
3=Good vs 5=Poor				
Intercept	204.441	97.897	426.940	
1=Male	1.167	0.928	1.467	
2=Female	1.000	•	•	
Current Smoker?				
1=Yes	0.644	0.469	0.885	
2=No	1.000			
AGE	0.948	0.942	0.954	
Weight (lbs)	0.999	0.996	1.003	;
4=Fair vs 5=Poor				
Intercept Sex	17.932	7.654	42.013	
1=Male	0.885	0.683	1.148	
2=Female	1.000			
Current Smoker?				
1=Yes	0.841	0.620	1.141	
2=No	1.000	•		
AGE	0.969	0.963	0.976	
Weight (lbs)	1.002	0.999	1.005	

Some interpretations from the table of default odds ratios in *Exhibit 8*: the odds of self-reporting Excellent or Very Good health are increased by 89% (OR=1.889) and 40% (OR=1.399), respectively, in males compared to females. Being a current smoker reduces the odds of self-reporting Excellent health by more than 70% (OR=0.282).

Variance Estimation N	Method: Taylor S	eries (WR) Using M	ultiply In	mputed Data	/
SE Method: Robust (Bi	inder, 1983)					
Working Correlations:	: Independent					
Link Function: Genera	alized Logit					
Response variable HAE	B1MI: Self-Ratin	g Health a	Status			
For Subpopulation: Ag	ge 20+					
Modelling Self-Rating	g of Health Stat	us (MI Da [.]	ta)			
Results for Summary (Over All Imputat	ions				
Solf-Dating Woolth St						
Predicted Marginal	lacus		Lower	Unner		
#1	Predicted		95%	05%		
11 ±	Marginal	9 F	Timi+	JJ 0 Timit	T.Mara-0	P=value
	Maryinar				1.Marg-0	I Vaiue
1=Excellent						
Sex						
1=Male	0.2477	0.0089	0.2302	0.2660	27.78	0.0000
2=Female	0.1760	0.0073	0.1618	0.1912	24.11	0.0000
2=Verv Good						
Sex						
1=Male	0.3120	0.0077	0.2966	0.3278	40.29	0.0000
2=Female	0.2953	0.0093	0.2769	0.3144	31.67	0.0000
3=Good						
Sex						
1=Male	0.3055	0.0073	0.2910	0.3205	41.70	0.0000
2=Female	0.3406	0.0093	0.3221	0.3596	36.54	0.0000
4=Fair						
Sex						
1=Male	0.1054	0.0055	0.0949	0.1169	19.33	0.0000
2=Female	0.1514	0.0079	0.1363	0.1679	19.28	0.0000
5=Poor						
Sex						
	0 0204	0 0023	0.0252	0.0343	13.07	0.0000
1=Male	0.02.94	0.002.7			/	0 0

Exhibit 9 presents the predicted marginal proportions (model-adjusted risks) and their 95% confidence limits for each category of health status, separately for each sex. It becomes clear that men are more likely to report being in Excellent health than are women (25% vs. 18%, respectively, with confidence limits that do not overlap). Women are more likely to report being in the less desirable categories of Good, Fair, or Poor health compared to men. This is after adjusting for age, weight, and current smoking status.

Exhibit 10. Model-Adjusted Risk Differences

Variance Estimation Method: Taylor Series (WR) Using Multiply Imputed Data SE Method: Robust (Binder, 1983) Working Correlations: Independent Link Function: Generalized Logit Response variable HAB1MI: Self-Rating Health Status For Subpopulation: Age 20+									
Modelling Self-Rating of Heal	th Status (MI I	Data)							
Results for Summary Over All	Imputations								
Self-Rating Health Status Contrasted Predicted Marginal #1	PREDMARG Contrast	SE	T-Stat	P-value					
1=Excellent Males-Females 2=Very Good Males-Females 3=Good Males-Females 4=Fair Males-Females 5=Poor	0.0716 0.0166 -0.0351 -0.0460	0.0080 0.0103 0.0091 0.0070	8.94 1.61 -3.84 -6.54	0.0000 0.1140 0.0004 0.0000					
Males-Females	-0.0072	0.0033	-2.18	0.0344					
NHANES III, Multiply Imputed	Dataset: Respo	ondents Age	20+						

Exhibit 10 shows that the difference in male vs. female risk is significant for all but one category of response. Men are significantly more prevalent than women in the Excellent health category, while women are significantly more prevalent than men in the Good, Fair, and Poor health categories.

Exhibit 11. Model-Adjusted Risk Ratios

Variance Estimation Method: Ta SE Method: Robust (Binder, 198 Working Correlations: Independ Link Function: Generalized Log Response variable HABIMI: Sel: For Subpopulation: Age 20+	aylor Serie 33) dent git E-Rating He	es (WR) (ealth Sta	Jsing Mu. atus	ltiply In	puted Data	
Modelling Self-Rating of Healt	in Status	(MI Data))			
Results for Summary Over All 3	Imputations	3				
Self-Rating Health Status Predicted Marginal Risk Ratio #1	Risk Ratio	SE	Lower 95% Limit	Upper 95% Limit		
1=Excellent Sex						
1=Male vs. 2=Female 2=Very Good Sex	1.407	0.053	1.303	1.519		
1=Male vs. 2=Female 3=Good Sex	1.056	0.036	0.986	1.132		
1=Male vs. 2=Female 4=Fair Sex	0.897	0.025	0.848	0.949		
1=Male vs. 2=Female 5=Poor Sex	0.696	0.036	0.627	0.773		
1=Male vs. 2=Female	0.804	0.081	0.657	0.984		
NHANES III, Multiply Imputed I	Dataset: I	Responder	nts Age 1	 20+		

The ratio of male vs. female risk is presented in *Exhibit 11*, above. The risk of self-reported Excellent health is increased by 41% (RR=1.407) for men vs. women. Conversely, the risk of being in Good, Fair, or Poor health is reduced by 10%-30% in men vs. women (RR=0.897, 0.696, and 0.804 for Good, Fair, and Poor health, respectively). All but one of the 95% confidence limits do not contain the null value of 1.0.